

 <b>aptara</b> <small>The Content Transformation Company</small>	PEPS	peps_1237	<b>Dispatch:</b> December 5, 2011	<b>CE:</b> AFL
	<b>Journal</b>	<b>MSP No.</b>	<b>No. of pages:</b> 48	<b>PE:</b> Jon

1 PERSONNEL PSYCHOLOGY  
2 2012, 65, 1–48

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5 **WHY EMPLOYEES DO BAD THINGS: MORAL**  
6 **DISENGAGEMENT AND UNETHICAL**  
7 **ORGANIZATIONAL BEHAVIOR**  
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24 We examine the influence of individuals' propensity to morally disengage on a broad range of unethical organizational behaviors. First, we develop a parsimonious, adult-oriented, valid and reliable measure of an individual's propensity to morally disengage, and demonstrate the relationship between it and a number of theoretically relevant constructs in its nomological network. Then, in four additional studies spanning laboratory and field settings, we demonstrate the power of the propensity to moral disengage to predict multiple types of unethical organizational behavior. In these studies we demonstrate that the propensity to morally disengage predicts several outcomes (self-reported unethical behavior, a decision to commit fraud, a self-serving decision in the workplace, and co-worker- and supervisor-reported unethical work behaviors) beyond other established individual difference antecedents of unethical organizational behavior, as well as the most closely related extant measure of the construct. We conclude that scholars and practitioners seeking to understand a broad range of undesirable workplace behaviors can benefit from taking an individual's propensity to morally disengage into account. Implications for theory, research, and practice are discussed.

39  
40 A host of ethical debacles across a wide range of contexts has inspired growing interest in studying and understanding why individuals engage in the kind of behavior that leads to enormous costs—trillions

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of dollars annually—for organizations and society. As just one example, the Association of Certified Fraud Examiners recently estimated that businesses globally suffer annual losses of \$2.9 trillion as a result of fraudulent activity (2010). This is a huge sum, indicating that unethical behavior is far more widespread than suggested by the intense focus on the few high-profile scandals that are covered by major news outlets. Thus, being able to better understand and predict who is likely to engage in such behavior—that is, behavior within organizations that directly cause direct harm to another individual or that violate widely accepted moral norms in society—is crucial for organizational leaders, and for societal well-being.

Organizational scholars have begun to identify a number of important contextual drivers of unethical organizational behavior, such as ethical leadership (Brown & Treviño, 2006), ethical climate (Mayer, Kuenzi, & Greenbaum, 2009), and codes of conduct (Weaver & Treviño, 1999). However, research thus far has failed to explain a substantial proportion of the variance in unethical organizational behavior using contextual variables alone (Kish-Gephart, Harrison, & Treviño, 2010). A range of individual-level factors has also been used to aid in the explanation of why people engage in unethical organizational behavior. The list of these antecedents is long, including Machiavellianism (Christie & Geis, 1970; Shultz, 1993; Siegel, 1973), moral identity (Aquino & Reed, 2002; McFerran, Aquino, & Duffy, 2010), cognitive moral development (Ambrose, Arnaud, & Schminke, 2008; Greenberg, 2002; Kohlberg, 1969), moral philosophies (Bird, 1996; Forsyth, 1980), empathy (Davis, 1983; Gino & Pierce, 2009), and moral affect (Eisenberg, 2000; Paternoster & Simpson, 1996; Tangney, 1990). However, effect sizes for their role in predicting unethical workplace behavior are generally small, leaving much variance unexplained (Kish-Gephart et al., 2010).

In this paper, we propose that an important additional driver of unethical behavior is an individual's propensity to morally disengage—that is, an individual difference in the way that people cognitively process decisions and behavior with ethical import that allows those inclined to morally disengage to behave unethically without feeling distress (Bandura, 1990a, 1990b, 1999, 2002). Broadly speaking, we know that how individuals process, frame, or understand information relevant to ethically meaningful decisions plays an important role in their ethical and unethical choices (Kern & Chugh, 2009; Tenbrunsel & Messick, 1999), and recent reviews of the behavioral ethics literature have suggested that scholars should attend more carefully to the role of cognitive processes in unethical behavior (Tenbrunsel & Messick, 2004; Tenbrunsel & Smith-Crowe, 2008). In the set of studies reported here, we heed these calls by

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establishing a new measure of the propensity to morally disengage as a uniquely important predictor of a broad range of unethical behaviors. **Q1**

This research makes two major contributions. First, though there has been a surge of interest in the concept of moral disengagement in the past decade or so, researchers have not yet offered the field a carefully validated, parsimonious, and easily administered measure of the general propensity to morally disengage. To date, studies of moral disengagement have depended on idiosyncratic methods of assessment, with new measures employed for each study, often without systematic development (e.g., Detert, Treviño, & Sweitzer, 2008; McFerran et al., 2010) or without tapping the construct in a comprehensive way (Aquino, Reed, Thau, & Freeman, 2007). A number of scholars have designed measures with a specific audience in mind, such as children (Bandura, Barbaranelli, Caprara, & Pastorelli, 1996), athletes (Boardley & Kavussanu, 2007; Corrion, Scoffier, Gernigon, Cury, & d'Arripe-Longueville, 2010), computer hackers (Rogers, 2001), or racial minorities (Pelton, Gound, Forehand, & Brody, 2004), or with a specific outcome in mind, such as support for military force (McAlister, 2001), the death penalty (Osofsky, Bandura, & Zimbardo, 2005), or violating one's duties to civic society (Caprara, Fida, Vecchione, Tramontano, & Barbaranelli, 2009). Our first goal is thus to provide a measure that can be easily administered and used generally—that is, with any adult sample in any type of context (though we focus here on the workplace context)—to successfully predict a wide set of unethical behaviors. In doing so, we address a long-standing concern among organizational ethics scholars about the lack of valid and reliable scales to use in research (Mayer et al., 2009; Tenbrunsel & Smith-Crowe, 2008; Treviño, Weaver, & Reynolds, 2006).

Second, research on moral disengagement has often been conducted without an adequate understanding of its role within the existing landscape of individual predictors of unethical behavior. We thus propose how the propensity to morally disengage should relate to other individual difference constructs of three specific types: (1) morally relevant personality traits, (2) moral reasoning abilities and orientations, and (3) dispositional moral emotions. We empirically position the propensity to morally disengage within the landscape of these other constructs and demonstrate that, compared to them, it is a more powerful predictor of four different measures of unethical organizational behavior. Finally, given the strength of the propensity to morally disengage as a predictor of unethical behavior, and the fact that most studies with other individual difference predictors have not included a strong overarching theoretical framework, we suggest that researchers consider adopting Bandura's (1986) theory of self-regulation as a conceptual framework that may lead to better

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3 understanding of how these individual differences operate and collec-  
4 tively explain unethical behavior.

5 In what follows, we present the rationale and results of five studies that  
6 build on and complement one another. We used Study 1 to develop and  
7 establish the baseline psychometric properties of a parsimonious, general  
8 measure of the propensity to morally disengage. Then, in each of Studies  
9 2–5, we used an array of unethical behavior measures as dependent vari-  
10 ables (measured independently from the propensity to morally disengage  
11 in each case) to test the general hypothesis that the propensity to morally  
12 disengage will explain a significant proportion of the variance in uneth-  
13 ical organizational behavior after controlling for other variables within  
14 its nomological net. The alternate predictors included here are among the  
15 most theoretically relevant and among the most common individual differ-  
16 ences used in ethical decision-making studies. However, they are almost  
17 never studied together nor are they adequately controlled for (as we do  
18 here) when demonstrating the utility of an additional construct. As part  
19 of this endeavor, we also show the superior ability of our new measure  
20 to predict both supervisor and coworker reported employee unethical be-  
21 havior over the most relevant extant measure of the propensity to morally  
22 disengage in the workplace (McFerran et al., 2010). Together, these stud-  
23 ies offer robust evidence of the power of an individual's propensity to  
24 morally disengage to predict a host of unethical behaviors of interest to  
25 organizational scholars and leaders.

#### 26 27 28 Theoretical Background

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30 Albert Bandura introduced the theory of moral disengagement as an  
31 extension of his more general social cognitive theory (Bandura, 1986:  
32 pp. 375–389). According to social cognitive theory, when self-regulatory  
33 capabilities are working properly, transgressive behavior is deterred  
34 through the self-condemnation individuals anticipate they would suffer  
35 were they to engage in behavior that conflicts with their internalized moral  
36 standards. Moral disengagement theory explains how this self-regulatory  
37 process can fail when moral disengagement mechanisms disable the cog-  
38 nitive links between transgressive behavior and the self-sanctioning that  
39 should prevent it (Bandura, 1986: pp. 376–385, Bandura, 1990a, 1990b,  
40 1999, 2002). The moral disengagement process is theorized to play an  
41 important role in explaining how individuals are able to engage in human  
42 atrocities such as political and military violence (Bandura, 1990a, 1990b),  
43 or corporate wrongdoing and corruption (Bandura, Caprara, & Zsolnai,  
44 2000; Brief, Buttram, & Dukerich, 2001; Moore, 2008b) without apparent  
45 cognitive distress.

*Moral Disengagement Mechanisms*

Bandura proposed that moral disengagement occurs through a set of eight inter-related cognitive mechanisms that facilitate unethical behavior. Moral justification, euphemistic labeling, and advantageous comparison are three mechanisms of moral disengagement that serve to cognitively restructure unethical acts so that they appear less harmful. *Moral justification* cognitively reframes unethical acts as being in the service of a greater good. Illustrations include the justifying of military atrocities as serving a worthy goal (Kramer, 1990; Rapoport & Alexander, 1982), or the recasting of inappropriate behavior such as unfair treatment as appropriate to protect friends or an organization. *Euphemistic labeling* is the use of sanitized language to rename harmful actions to make them appear more benign (Bolinger, 1982). For example, in corrupt organizations, those who collude are often positively labeled “team players” (see Jackall, 1988, pp. 52–53). *Advantageous comparison* exploits the contrast between a behavior under consideration and an even more reprehensible behavior to make the former seem innocuous (Bandura, 2001). For example, misrepresenting small lies on expense reports can be viewed as more acceptable when compared with more egregious expense report violations.

The displacement and diffusion of responsibility mechanisms obscure the moral agency of the (potential) actor. *Displacement of responsibility* refers to the attribution of responsibility for one’s actions to authority figures who may have tacitly condoned or explicitly directed behavior (see Kelman & Hamilton, 1989; Milgram, 1974; Sykes & Matza, 1957). *Diffusion of responsibility* works in a similar way, but refers to dispersing responsibility for one’s action across members of a group (see the description of the lead-up to the Challenger disaster recounted by Vaughan, 1996).

Distortion of consequences, dehumanization, and the attribution of blame mechanisms serve to reduce or eliminate the distress one perceives to be causing a victim (Sykes & Matza, 1957). *Distortion of consequences* describes the minimization of the seriousness of the effects of one’s actions, thus providing “little reason for the self-censure to be activated” (Bandura, 1999b, p. 199). This is illustrated by descriptions of stealing from a large, profitable organization as a “victimless crime” (Benson, 1985). *Dehumanization* is the framing of the victims of one’s actions as undeserving of basic human consideration. This is fostered by defining others as members of an out-group who are unworthy of moral regard (Deutsch, 1990; Opatow, 1990). Finally, in *attribution of blame*, responsibility is assigned to the victims themselves, who are described as deserving whatever befalls them (Bandura, 2002, p. 110). It has been shown to describe the cognition underlying unethical behavior in many contexts, including types of white collar crime (see Douglas, 1995).

Although others have discussed or studied similar cognitive mechanisms separately (e.g., euphemistic language, diffusion of responsibility; see Ashforth & Anand, 2003; Diener, 1976; Kelman, 1973), Bandura conceptualized these eight moral disengagement mechanisms as a coherent set of cognitive tendencies that manifest in individuals as a trait, and influence the way individuals may approach decisions with ethical import. The trait approach—which has characterized most of Bandura’s work and represents our perspective here—argues that individuals will systematically differ in their propensities to use cognitive moral disengagement mechanisms when facing decisions with ethical import. In the set of studies reported here, we focus on the trait instantiation of moral disengagement, examining the propensity to morally disengage as a generalized cognitive orientation to the world that differentiates individuals’ thinking in a way that powerfully affects unethical behavior. This stance is consistent the majority of the empirical work on moral disengagement to date, from Bandura and colleagues’ (1996) study of the propensity to morally disengage in children and adolescents to a number of subsequent studies carried out by other researchers using a range of measures of the individual propensity to morally disengage (e.g., Boardley & Kavussanu, 2007; Caprara et al., 2009; Detert et al., 2008; Duffy, Tepper, & O’Leary-Kelly, 2002; McAlister, 2001).

#### *The Nomological Network of the Propensity to Morally Disengage*

A nomological network is a conceptual model that situates a construct of interest within the landscape of constructs that are theoretically related to it (Schwab, 1980). Though identifying and testing relationships within a nomological network is often narrowly understood as part of the process of construct validation and measurement development (Chronbach & Meehl, 1955), the work of specifying a construct’s nomological network for measurement purposes also helps researchers fully understand a construct’s theoretical implications. In this section, we describe the nomological network of the propensity to morally disengage and, in so doing, explain why the propensity to morally disengage is likely to be a particularly strong predictor of unethical organizational behavior relative to other constructs that share common conceptual space.

We identify three important categories of constructs within the nomological network of the propensity to morally disengage: (1) morally relevant individual personality traits, (2) moral reasoning abilities and orientations, and (3) dispositional moral emotions. For each category and construct, we explain the rationale for inclusion and the expected relationship between the nomological network factor, the propensity to moral disengage (our focal construct), and the key criterion variable (unethical

behavior). Our intention was not to exhaustively tap every possible category or construct of potential relevance to the propensity to morally disengage, but to choose theoretically salient and well-studied representative constructs within each of these three conceptual categories.

#### *Morally Relevant Individual Traits*

As noted earlier, our focus is on the dispositional propensity of individuals to morally disengage (Bandura, 1990a, 1990b). Thus, other stable dispositions that have been found to influence ethical and unethical behavior and that describe orientations toward or ways of seeing behaviors with ethical ramifications should fall within its nomological network. Three in particular have been consistently identified as important predictors of ethical and unethical behavior (Hoffman, 2000; Kish-Gephart et al., 2010): *Machiavellianism* (Christie & Geis, 1970), *moral identity* (Aquino & Reed, 2002), and *empathy* (Davis, 1983).

*Machiavellianism* represents an individual's propensity to be manipulative and ruthless in the pursuit of self-interested goals (Christie & Geis, 1970). We reason that those high in Machiavellianism will be more inclined to morally disengage because such cognitive mechanisms present one means by which Machiavellians can more readily pursue their own interests without self-censure. Machiavellianism has been shown to be positively related to many transgressive behavioral tendencies, including anti-social behavior, lying, and willingness to exploit others (Christie & Geis, 1970; Sakalaki, Richardson, & Thépaut, 2007), as well as to a broad range of unethical decisions in a recent meta-analysis (Kish-Gephart et al., 2010). However, we believe that the propensity to morally disengage may be a particularly strong predictor of unethical organizational behavior because it captures an individual's general tendency to disengage from the self-sanctions that would otherwise prevent a wide range of unethical behaviors, rather than the narrower and more specific behaviors associated with Machiavellianism.

*Moral identity* describes the extent to which one's self-concept incorporates the importance of being a moral person (Aquino & Reed, 2002). Moral identity has a strong relationship with pro-social behavior and has also been linked with reduced unethical behavior (Aquino & Reed, 2002; Shao, Aquino, & Freeman, 2008). We expect moral identity to be negatively correlated with the propensity to morally disengage because individuals with a highly salient moral identity should be more concerned about harm to others and more likely to take responsibility for their behavior, thereby making it more unlikely that they would disengage the moral self-regulatory function (c.f., Detert et al., 2008). Further, we believe that the propensity to morally disengage will have stronger effects on behavior

because moral identity's effect requires an activated moral self-concept (Aquino & Reed, 2002). However, many of the mechanisms of moral disengagement disrupt the activation of the self-concept by, for example, focusing on the victim (dehumanization, attribution of blame) or by reducing personal agency (diffusion and displacement of responsibility).

*Trait empathy*, which includes sympathetic feelings, responsiveness to others, and an ability to cognitively understand others' perspectives, has received particular attention as an individual difference that contributes to ethical behavior and reduces unethical behavior (Eisenberg, 1986; Eisenberg & Miller, 1987; Hoffman, 2000; Tangney, 1991). We expect trait empathy to be negatively related to the propensity to morally disengage because those predisposed to morally disengage should be less likely to take others' viewpoints or feel compassionate towards them (Dertert et al., 2008). Thus, those lower in trait empathy (and thus less likely to feel compassionately towards others) will likely demonstrate higher propensities to morally disengage because the latter often also involves ignoring or distorting others' feelings, needs, or perspective. However, compared to trait empathy, morally disengaged reasoning is relevant to a much broader set of situations in work organizations. Therefore, we expect the propensity to morally disengage to have even greater general utility in the prediction of unethical organizational behavior.

#### *Moral Reasoning Abilities and Orientations*

The ethical decision making literature has traditionally been dominated by rational/deliberative models of moral reasoning, which suggest that individuals move from awareness to deliberative judgment to motivation/intention and then to action (Rest, 1986). Constructs describing aspects of reasoning about ethical issues should relate to the propensity to morally disengage because the disengagement process is also inherently cognitive and influences, probably less consciously, the framing and making of ethically charged decisions. The three main constructs that represent different deliberative modes of ethical reasoning are *cognitive moral development*, *idealism*, and *relativism*.

*Cognitive moral development* (CMD) is considered theoretically important to the judgment phase of ethical decision making—that is, the point when an individual decides what is right or wrong in a particular situation. CMD is conceptualized in terms of a series of stages through which individuals progress as they become more cognitively advanced and autonomous in their moral reasoning (Kohlberg, 1969, 1984; Piaget, 1965). Higher levels of CMD have been found to be negatively associated with unethical choice (see Kish-Gephart et al., 2010), but the construct's potential relationship to the overall moral disengagement process or



specific mechanisms of disengagement has never been studied. We argue that CMD is conceptually distinct from the propensity to morally disengage because the former is a measure of the level of sophistication with which one consciously reasons through ethical quandaries, whereas the propensity to morally disengage describes a dispositional tendency to use cognitive mechanisms that disengage moral self-regulatory sanctions. We propose that CMD will be negatively related to the propensity to morally disengage because those who engage in more complex, principled reasoning are more likely to quickly invoke moral principles (e.g., justice, the greater good) to evaluate ethical dilemmas and to decide what is right. They are also more likely to think through ethical dilemmas autonomously rather than relying on others for guidance. Therefore, they should also be less inclined to displace or diffuse responsibility onto others.

Despite its prominence in theory and research for decades, a recent meta-analysis reported that CMD is only moderately predictive of unethical choices (Kish-Gephart et al., 2010). This may be because the ability to reason deliberatively and in a sophisticated way about ethical decisions does not always lead one to behave more ethically (Rest, 1986). One possible explanation for this finding is that many unethical acts are better explained by impulsive or intuitive models (Haidt, 2001) rather than the deliberative model assumed by CMD. Although understanding of the overall process of moral disengagement is still unfolding, moral disengagement appears to often involve little conscious deliberation. Thus, we would expect not only a negative correlation between the propensity to morally disengage and CMD, but also that moral disengagement is potentially “in play” in a wider variety of situations (where conscious deliberation is not involved) and therefore will show a stronger relationship with unethical behavior than CMD.

It is also important to examine the relationship between the propensity to morally disengage and the most common ethical philosophies that guide human decision making in ethical dilemma situations. Ethical philosophies describe “stated beliefs or personal preferences for particular normative frameworks” (Kish-Gephart et al., 2010, p. 3). *Idealism* and *relativism* describe two moral philosophies that reflect stable individual orientations toward ethical decision making. Forsyth (1980; Schlenker & Forsyth, 1977) described idealism as an individual’s belief that “the ‘right’ action [can] always be obtained” and relativism as the degree to which an individual “rejects universal moral rules . . . when drawing conclusions about moral questions” (Forsyth, 1980: pp. 175–176). We reason that the propensity to morally disengage will be positively correlated with relativism because holding a relativist position is facilitated by morally disengaged cognitions. Conversely, the propensity to morally disengage should be negatively correlated with idealism because idealists are driven

to pursue absolute ethical standards and thus not be motivated to find ways (cognitively) to skirt them. Also, similar to cognitive moral development, idealism and relativism are part of a rational/deliberative approach to ethical decision making and would only be involved in situations involving conscious deliberation. As a result, we expect the propensity to morally disengage to apply across a wider range of situations including those that do not involve conscious ethical deliberation.

### *Dispositional Moral Emotions*

Because moral disengagement facilitates engaging in unethical behavior without feeling distress (Bandura, 1990a) moral affect/emotion constructs should also fall within the nomological network of the propensity to morally disengage. Moral emotions, including anticipatory guilt and shame, have been recognized as important motivating factors underlying ethical conduct (Haidt, 2003; Tangney, Stuewig, & Mashek, 2007) but they have never been studied in relation to moral disengagement. The propensity to morally disengage is conceptually distinct from moral affect because the former describes a set of relatively “cool” cognitive processes stemming from parts of the brain not primarily involved in the “hot” processing included in moral affect. However, as described below, we expect that the propensity to morally disengage will be related to certain dispositional moral emotions.

Guilt has been well established as a correlate of ethical behavior (Tangney et al., 2007, p. 354), in part because it elicits a sense of personal responsibility for one’s actions (Tangney, 1991). We expect the propensity to morally disengage to be negatively related to *dispositional guilt*—the tendency to experience a set of negative emotions upon judging one’s own actions as harmful or immoral (Lewis, 1971; Tangney, 1991)—because guilt is activated when one’s self-sanctions against unethical behavior are working correctly, and morally disengaged reasoning weakens that self-sanction trigger. Shame, however, works differently. It is a set of negative emotions about one’s self rather than one’s actions (Lewis, 1971; Tangney, 1991). Because the propensity to morally disengage represents a tendency to use certain types of cognitions to distance oneself from one’s actions (and thus not see behavior in specific situations as being “reflective of” or “revealing of” the self), we expect the propensity to morally disengage to be unrelated to *dispositional shame*. As noted above, an individual’s self-sanctions against unethical behavior must be working correctly in order for guilt to be activated. By contrast, the propensity to morally disengage can operate across a wide array of situations and assumes that self-sanctions may be disengaged. Therefore, we expect that the propensity

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3 to morally disengage will be more consistently related to unethical orga-  
4 nizational behavior than will guilt.

5  
6 *Hypotheses*  
7

8 Thus far, we have provided theoretical rationale for expected relation-  
9 ships between the propensity to morally disengage and other constructs  
10 within its nomological network as well as its distinctiveness from these  
11 other constructs. However, our ultimate interest is in better understanding  
12 what explains organizationally embedded unethical behavior—behavior  
13 that is widespread enough to lead to the huge annual financial losses  
14 mentioned at the outset of this paper. Thus, we turn next to the rationale  
15 for our expectation that the propensity to morally disengage is a dis-  
16 tinct and powerful predictor of the types of unethical behavior relevant to  
17 organizations.

18 The workplace provides ample opportunities for moral disengage-  
19 ment: organizations tend to be hierarchical, providing opportunities for  
20 the displacement of responsibility; work is often undertaken within teams,  
21 providing opportunities for the diffusion of responsibility; organizational  
22 membership automatically defines the boundaries of an in-group, pro-  
23 viding opportunities for moral justification (to protect the organization)  
24 and the cognitive minimization of the consequences of one's actions for  
25 those who are outside the organization (and thus in an out-group). The  
26 propensity to morally disengage might also be particularly damaging  
27 in organizational life because work contexts have been documented as  
28 triggering amoral frames of judgment (Gross, 1978). As Jackall pointed  
29 out in *Moral Mazes*, organizations are particularly effective at assist-  
30 ing individuals in bracketing off moral schemas that guide behavior  
31 elsewhere (Jackall, 1988). Thus, the propensity to morally disengage is  
32 likely to be particularly relevant in the prediction of unethical behavior in  
33 organizations.

34 Though the link has been suggested previously (Bandura et al., 2000;  
35 Moore, 2008b), little research on the negative outcomes of moral dis-  
36 engagement extends directly to unethical behavior in organizations (see  
37 Duffy, Aquino, Tepper, Reed, & O'Leary-Kelly, 2005, for an exception).  
38 Among the few empirical studies to date that have documented specific  
39 negative outcomes of the propensity to morally disengage, a couple link  
40 this propensity to organizationally relevant generic behaviors such as  
41 cheating, lying, and stealing (Detert et al., 2008). However, no study has  
42 yet shown that a carefully validated measure of the propensity to morally  
43 disengage can predict organizationally relevant unethical behavior or un-  
44 ethical behavior among employed adults, after controlling for a full array  
45 of other constructs likely to be related strongly to either the propensity to

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3 morally disengage itself, or to unethical behavior. Thus, we propose and  
4 test here not merely that the propensity to morally disengage is distinct  
5 from the other types of morally salient individual differences discussed  
6 above, but that this propensity will predict a variety of types of unethical  
7 behavior after controlling for these alternative explanations.

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9 *Hypothesis 1:* The propensity to morally disengage will be positively  
10 related to unethical behavior after controlling for other  
11 morally salient individual traits.

12 *Hypothesis 2:* The propensity to morally disengage will be positively  
13 related to unethical behavior after controlling for de-  
14 liberative moral reasoning capacity and moral philoso-  
15 phies.

16 *Hypothesis 3:* The propensity to morally disengage will be positively  
17 related to unethical behavior after controlling for dis-  
18 positional moral emotions.

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20 Beyond establishing that the propensity to morally disengage is a  
21 non-redundant predictor of unethical behavior, when introducing a new  
22 measure it is also important to demonstrate its value relative to existing  
23 measures of the same construct. As noted earlier, most of the few extant  
24 measures of the propensity to morally disengage are not clearly geared  
25 toward adults, or designed to tap all the mechanisms of the propensity  
26 to morally disengage. The measure closest to achieving all of these aims  
27 is the one developed by Duffy and colleagues (Duffy et al., 2005; Duffy  
28 et al., 2002; McFerran et al., 2010). This measure was developed to explain  
29 undesirable and unethical behaviors at work, such as social undermining  
30 of colleagues and organizational deviance (Duffy et al., 2005; Duffy et al.,  
31 2002). Thus, we also examine how our new measure of the propensity to  
32 morally disengage compares to this existing measure, including whether  
33 the new measure still predicts unethical behavior in a workplace context  
34 after accounting for the variance explained by the alternative measure.

35 *Hypothesis 4:* The new measure of the propensity to morally disen-  
36 gage offered here will be positively related to unethi-  
37 cal employee behavior after controlling for an existing  
38 measure of the propensity to morally disengage.

#### 40 Measure Development and Assessment

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42 Our development and validation of a new measure of the propensity to  
43 morally disengage was undertaken to address several specific limitations  
44 of existing measures. First, the original measure developed by Bandura  
45 and colleagues (1996) was developed specifically for use with children and

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3 is therefore not appropriate for adults—items, for example, include refer-  
4 ence to schoolyard pranks and classroom teasing. Second, we also wanted  
5 to develop a measure that incorporates all of the mechanisms of moral  
6 disengagement rather than one or a few mechanisms (e.g., moral justifi-  
7 cation; see Aquino et al., 2007). Third, we wanted a measure that would  
8 be appropriate for a broad sample of adults. Other measures have been  
9 developed for a particular type of sample or context, which significantly  
10 limits their general use (e.g., Boardley & Kavussanu, 2007; Caprara et al.,  
11 2009; Corrion et al., 2010; McAlister, 2001; Pelton et al., 2004; Rogers,  
12 2001). Fourth, we aimed to provide the first systematic documentation  
13 of the convergent, discriminant, and predictive validity (across a wide  
14 range of samples) of a measure of the propensity to morally disengage  
15 (e.g., Detert et al., 2008; McFerran et al., 2010). The final version of the  
16 measure we introduce here is also significantly more parsimonious (only  
17 eight items) than existing measures (which range from 15 to 34 items  
18 (e.g., Bandura et al., 1996; Detert et al., 2008; Duffy et al., 2005), making  
19 it highly advantageous for use in field research.

20 Consistent with Bandura’s theoretical claim that moral disengagement  
21 is best understood to be “multifaceted” (Bandura et al., 1996, p. 367), not  
22 multi-factorial, and in line with both his (e.g., Bandura et al., 1996) as  
23 well as subsequent published and unpublished work on moral disengage-  
24 ment (Detert et al., 2008; Duffy et al., 2005; Moore, 2008a), our aim was  
25 to create a unidimensional measure of the general propensity to morally  
26 disengage. That is, while acknowledging that the eight individual mecha-  
27 nisms of moral disengagement represent different facets of the construct,  
28 our overarching goal was to tap these facets as part of a valid scale that  
29 assesses the general propensity to morally disengage as a higher order  
30 concept.

31 We began by pre-testing a large pool of items (74) in a sample of 454  
32 full-time employed adults (26% male,  $M_{\text{age}} = 36.3$  years,  $SD = 9.8$  years)  
33 to ensure ample representation of the construct’s broad content domain  
34 and to assess the properties of various items when rated by a wide range  
35 of adults. We chose items to represent each of the mechanisms of moral  
36 disengagement, to be understandable to adults across multiple contexts  
37 and cultures, and to represent cognitions about general behaviors (such  
38 as lying, cheating, or stealing)—thus measuring a *general* propensity to  
39 morally disengage—rather than behaviors in specific contexts that might  
40 not be relevant to all adults (e.g., lying in a real estate negotiation). In de-  
41 veloping these items we were guided by Bandura’s theoretical descriptions  
42 of the eight moral disengagement mechanisms (Bandura, 1990a, 1990b,  
43 1999, 2002), as well as by previous measures (Bandura et al., 1996;  
44 Detert et al., 2008; Duffy et al., 2005; Moore, 2008a). We wrote new items  
45 when we judged that existing items for a particular moral disengagement

mechanism might not meet our criteria for developing a general scale appropriate for most adults. We used the pre-test to select items with the highest factor loadings on a single factor, and with the widest and least skewed distributions. The pre-test left us with 47 items to further assess in Study 1.

### *Study 1*

Study 1 was conducted to further examine the psychometric properties of the set of items developed in the pre-test, and to determine how parsimoniously the propensity to morally disengage could be measured while still ensuring that the multi-faceted nature of the construct was represented by the scale. Prior to making decisions about scale length, it was also necessary to thoroughly examine the dimensionality of the construct—that is, to ensure that moving forward with a unidimensional measure was empirically justified. Having done this, we used statistical indices and researcher judgment to select items that created three increasingly shorter measures of the propensity to morally disengage (a 24-item, 16-item, and 8-item version). We then compared these three scales' psychometric properties and preliminary validity evidence, with the goal of determining the most parsimonious way the construct could be measured without unacceptably compromising on reliability or validity standards.

### *Methods and Measures*

*Participants and Procedure.* One hundred and ninety four adults, who were recruited to participate through a university-based behavioral lab in the United Kingdom, completed a web-based survey including measures of some of the variables within the construct's nomological network. The sample was diverse in age ( $M_{\text{age}} = 25.6$ ,  $SD = 6.7$ ) and background (see Table 1 for demographic details). One hundred and sixteen of these participants also participated subsequently in a second study in the lab. Participants were paid £5 for completing the web-based survey, and £10 for participating in the lab study.

*The Propensity to Morally Disengage.* The 47 items developed in the pre-test were used to create and compare potential versions of a propensity to morally disengage measure: a 24-item version (3 items per moral disengagement mechanism), a 16-item version (2 items per mechanism), and an 8-item version (1 item per mechanism). Appendix A lists all the items, and specifies that are included in each version of the scale. Because our aim was to create a measure that represents a higher-order concept tapping each of the eight specific mechanisms of moral disengagement, we ran a factor analysis model forcing all items onto a single factor.

TABLE 1  
*Overview of Samples Used in Studies 1–5, with Descriptive Statistics*

	Gender		Age		Race/Background
	Female	Male	<i>M</i>	<i>SD</i>	
Study 1: U.K.-based University Behavioral Lab Participants ( <i>n</i> = 194)	114	80	25.6	6.7	White = 383, African American = 17, Asian = 26, Hispanic = 18, other/declined to answer = 10
Study 2: Northeastern U.S. undergraduates ( <i>n</i> = 272)	109	163	19.1	0.6	White = 228, African American = 5, Asian = 14, Hispanic = 5, foreign country native = 12, other/declined to answer = 7
Study 3: International MBA Students ( <i>n</i> = 304)	78	226	28.5	2.2	North America = 61, Europe = 107, Asia = 67, Latin/South America = 39, Middle East = 14, Africa = 8, Australia/New Zealand = 8
Study 4: Northeastern U.S. undergraduates ( <i>n</i> = 250)	105	145	18.9	1.0	White = 187, African American = 5, Asian = 25, Hispanic = 9, foreign country native = 18, other/declined to answer = 6
Study 5: Employees in the Southern U.S. ( <i>n</i> = 123)	69	67 <sup>1</sup>	25.7	9.8	White = 87, Hispanic or Latino/a = 26, African American = 15, Asian = 6, biracial = 2, other/declined to answer = 5

<sup>1</sup>Five respondents declined to state their gender.

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2  
3 The three highest-loading items representing each moral disengagement  
4 mechanism were selected as a 24-item measure of moral disengagement  
5 ( $\alpha = .90$ ). This 24-item measure was then trimmed to 16 items by dropping  
6 one item for each moral disengagement mechanism ( $\alpha = .88$ ). Items were  
7 dropped based on a combination of statistical (e.g., factor loadings) and  
8 theoretical (e.g., ensuring each tactic was most clearly represented by the  
9 items) considerations. Finally, the scale was trimmed to eight items ( $\alpha =$   
10  $.80$ ), with each moral disengagement mechanism represented by a single  
11 item. The choice of which items to drop was in this case based primarily  
12 on theoretical grounds, attempting to keep a highly representative item for  
13 each mechanism, while ensuring that the measure still had broad content  
14 coverage and acceptable estimated reliability.

15 *Convergent and discriminant validity measures.* Machiavellianism  
16 was measured using the standard 20-item Mach IV (Christie, 1970), which  
17 asks respondents to rate their level of agreement with a series of statements  
18 (on a 5-point continuum from “strongly disagree” to “strongly agree”).  
19 A sample item is, “It is hard to get ahead without cutting corners here  
20 and there.” Estimated reliability of the measure in this sample ( $\alpha =$   
21  $.65$ ) was low (though low reliabilities in the Mach IV are typical, see  
22 Ray, 1983; Wrightsman, 1991). The measure of moral identity (Aquino  
23 & Reed, 2002) presents a set of characteristics of “moral” people (e.g.,  
24 caring, compassionate, fair) as stimuli and then asks respondents (using  
25 a 7-point continuum from “strongly disagree” to “strongly agree”) to  
26 assess how important it is to be viewed as an individual who shares  
27 those characteristics. We used the five items that assess “internalization”  
28 of a moral identity because this component measures the strength of  
29 individuals’ self-concept as moral people. A sample item (rated after  
30 seeing the list of ‘moral characteristics’) is, “I strongly desire to have these  
31 characteristics.” Estimated reliability was  $.87$ . We assessed empathy using  
32 two components of Davis’ (1983) Interpersonal Reactivity Index (IRI):  
33 a seven-item measure of perspective-taking (PT), or the ability to adopt  
34 others’ viewpoints, and a seven-item measure of empathetic concern (EC),  
35 or the tendency to feel compassion towards others. Respondents reported  
36 the degree to which they feel the items are true about themselves, from “not  
37 at all true” (1) to “very true” (7). A sample item in the perspective-taking  
38 subscale ( $\alpha = .75$ ) is, “I try to look at everybody’s side of a disagreement  
39 when I make a decision,” and in the EC subscale ( $\alpha = .78$ ) is, “I am often  
40 quite touched by things that I see happen.”

41 Because social desirability can affect reporting in studies of ethics-  
42 related beliefs and behaviors, we wanted to ensure that our new measure  
43 of the propensity to morally disengage was not highly correlated with  
44 a measure of social desirability bias. Thus all participants completed  
45 a 10-item version of the traditional 33-item Marlowe-Crowne measure



TABLE 2  
*Summary of Fit Indices for all CFA Models, Study 1*

	$\chi^2$	<i>df</i>	$\chi^2/df$	RMSEA	SRMR	NNFI	CFI
<b>24 item measures</b>							
24 items on 1 factor	588	252	2.54	.090	.080	.90	.91
24 items on 8 factors	507	224	2.26	.089	.078	.90	.91
24 items on 8 factors on 1 second-order	560	244	2.30	.092	.081	.91	.92
<b>16 item measures</b>							
16 items on 1 factor	254	104	2.44	.099	.075	.91	.93
16 items on 8 factors	193	76	2.54	.098	.070	.91	.94
16 items on 8 factors on 1 second-order	244	96	2.54	.100	.074	.91	.93
<b>8 item measure</b>							
8 items on 1 factor	27	20	1.35	.045	.043	.98	.99

of social desirability (Crowne & Marlowe, 1960; Strahan & Gerbasi, 1972). For this measure, participants responded to a series of yes/no questions about negative behaviors or cognitions that are thought to be universal but socially embarrassing or unattractive (e.g., “I never resent being asked to return a favor”), and therefore potentially elicit socially desirable responses. The number of times that a participant answers “no” is summed as a measure of his/her social desirability response bias.

### Results

Prior to comparing results for three single-factor measures of differing length, we first thoroughly explored the dimensionality of these three measures of the propensity to morally disengage in a multi-step process. We began by estimating a series of confirmatory factor analysis models using LISREL 8.8 (Jöreskog & Sörbom, 2009) for the two longer versions of the scale: one that forced all the items to load onto a single latent variable (a 1-factor model), one that forced the items to load onto eight different factors representing each of the moral disengagement mechanisms (an 8-factor model), and one model where eight first-order factors (comprised by forcing the items tapping each moral disengagement mechanism onto separate factors) were forced to load onto a second-order latent variable (the second-order factor model). The resultant fit indices for each of these models is presented in Table 2. The first thing the table makes clear is that all of the models fit the data adequately, with RMSEA values less than .10, SRMR values less than .08, and CFI and NNFI indices all greater than .90. In particular, the fit indices for a single-factor model run using the 8-item version of the scale indicate a good fit of the data to the model ( $\chi^2 = 26$ ,  $df = 20$ ,  $ns$ , NNFI = .98, CFI = .99, SRMR = .04,

RMSEA = .05) according to current standards—e.g., values of less than .05 for the RMSEA (Browne & Cudeck, 1993; Hu & Bentler, 1999) and CFI and NNFI values greater than .95 (Hu & Bentler, 1999).

The second thing evident from the factor analyses results summarized in Table 2 is that statistical criteria alone do not provide strong evidence in this case for preferring one factor structure over another. For the 16- and 24-item versions of the scale, the CFA models that fit the data to eight separate factors have very slightly better fit indices (only models with the same observed variables are directly comparable). This is an expected result given the two or three items assessing each specific mechanism are intended to be more inter-related with each other than with those items assessing the other mechanisms. Conversely, the fit indices for the models that represent the best match to Bandura's theoretical expectation that moral disengagement should be conceived of as "multifaceted" (Bandura et al., 1996, p. 367)—that is, as a higher-order construct that encompasses eight different mechanisms in a unified way—are slightly lower than the other models run with the same number of items. This slight decrement in fit (compared to models involving only a first-order factor) is expected mathematically (Marsh, 1987) and thus does not necessarily have practical import. The simplest latent variable models for the 16- and 24-item versions of the scale—those that load all items onto a single first-order latent factor—have fit indices in between those for the 8-factor and second-order factor models. However, differences in fit on the various indices are of an average magnitude of just .01, a size considered to indicate measurement invariance between comparable models (Cheung & Rensvold, 2002). With this much similarity in fit among the comparable models, along with Bandura's theorization, the rule of parsimony suggests selecting a single-factor structure for this new measure of the propensity to morally disengage (Kline, 1988).

To further assess the appropriateness of proceeding with a single-factor solution (i.e., a unidimensional approach), we then examined whether measuring the propensity to morally disengage in various multidimensional and unidimensional ways resulted in any substantive empirical differences. Specifically, we created several measures of the propensity to morally disengage that were either a) latent variable measures derived from the factor loadings from the various first-order or second-order CFA models or b) simple average measures derived by taking the average of the items forming the scale. The latent variable approach allows for differential weighing of each item and of each moral disengagement mechanism (in the second-order model), whereas the simple average approach (where each item is assigned equal weight) assumes a unifactorial structure of the data with all items representing the single underlying factor equally well. The three versions of the 24-item measure—one an average of the

scale items and the other two latent variable measures derived from the factor loadings for a single factor model, and the hierarchical second-order factor model, respectively—are all correlated with each other at  $r = .98$  or higher.<sup>1</sup> The three parallel versions (simple average and first- and second-order latent variables) created using 16 items, and the two parallel versions (simple average and first-order latent variable) created using eight items likewise have bivariate correlations all approaching 1.0. These results also suggest that there is no practical difference, at any of the three scale lengths, between the simplest item-average unidimensional measures and those that take into account differences in the weights assigned to particular moral disengagement items or mechanisms in the formation of latent variables.

Having concluded on statistical and practical grounds that measuring the propensity to morally disengage in more complex ways produces no meaningful advantage, we proceeded in a third step to compare different length (24-, 16- and 8-item) versions of the propensity to morally disengage scale constructed as simple averages of the items. As shown in Table 3, the intercorrelations among the three different-length versions of the scale are all above .90. Moreover, the correlations of the three different-length versions of the propensity to morally disengage scales with the other constructs assessed in Study 1 are strongly consistent with theory and predictions. For example, the 8-item propensity to morally disengage measure correlates positively with Machiavellianism ( $r = .44$ ,  $p < .01$ ), and negatively with moral identity ( $r = -.55$ ,  $p < .01$ ) and two facets of empathy: perspective taking ( $r = -.40$ ,  $p < .01$ ) and EC ( $r = -.46$ ,  $p < .01$ ). This measure is also uncorrelated with social desirability ( $r = .05$ , *ns*). All of these relationships are in the expected direction, and none is so strong as to suggest that the propensity to morally disengage is redundant with any of the other constructs. And, importantly, the correlations between the 8-, 16- and 24-item versions of the propensity to morally disengage measure and all of the nomological network variables assessed in Study 1 are substantively equivalent—that is, all are in the hypothesized directions and of similar magnitude. Thus, we determined that we could meet our objective of creating a more parsimonious measure of the propensity to morally disengage than has previously been offered

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<sup>1</sup>We did not include in these comparisons (for the 24- and 16-item scale versions) eight first-order latent variables representing each moral disengagement mechanism separately for two reasons: (1) the second-order latent variable is essentially doing this in the creation of the eight first-order latent variables subsequently loaded onto the second-order latent variable, and this approach better represents the theory of the propensity to morally disengage as a single higher order construct with subdimensions; (2) reliabilities for the two- and three-item scales representing each moral disengagement mechanism that result from this approach are systematically far below the conventional cutoff of .70.

TABLE 3  
*Means, Standard Deviations, Correlations among Study 1 Variables*

Variable	<i>M</i>	<i>SD</i>	1	2	3	4	5	6	7
1. Propensity to morally disengage (8-items) <sup>a</sup>	2.57	.99	(.80)						
2. Propensity to morally disengage (16-items) <sup>a</sup>	2.88	.92	.93**	(.88)					
3. Propensity to morally disengage (24-items) <sup>a</sup>	2.99	.88	.90**	.98**	(.90)				
4. Machiavellianism <sup>a</sup>	2.83	.42	.44**	.52**	.54**	(.65)			
5. Moral identity <sup>a</sup>	6.13	1.01	-.55**	-.54**	-.54**	-.34**	(.87)		
6. Perspective taking <sup>a</sup>	4.80	.91	-.40**	-.36**	-.38**	-.29**	.38**	(.75)	
7. EC <sup>a</sup>	5.22	.91	-.46**	-.46**	-.47**	-.31**	.56**	.38**	(.78)
8. Social desirability <sup>a</sup>	3.65	1.54	.05	.00	-.01	.04	-.15*	-.16*	.00

<sup>a</sup>*N* = 194; <sup>b</sup>*N* = 116. \*\**p* < .01. \**p* < .05. Where appropriate, alpha reliabilities appear along the diagonal.

in the literature by proceeding with the 8-item version of the scale (which has estimated reliability of .80 in Study 1).

To statistically confirm that our new 8-item measure of the propensity to morally disengage is not redundant with related constructs, we followed recommendations by Schwab (1980) and DeVellis (1991) in considering results from several confirmatory factor analysis models estimated using the eight items from the propensity to morally disengage measure as well as all items assessing the other nomological network constructs included in Study 1—namely, Machiavellianism, moral identity, perspective taking, and EC. We compared fit indices for the hypothesized model—that is, a 5-factor model where we forced all the items to load onto factors representing their theorized construct—with those for several 4-factor models where we forced the propensity to morally disengage items to load on the same factor as the items from one of the related constructs (i.e., a 4-factor model forcing the moral disengagement items to load with the items for Machiavellianism, a 4-factor model forcing the moral disengagement items to load with the moral identity items, etc.). If any of the 4-factor models fit the data significantly better than the 5-factor model, it would indicate that our new measure may be redundant with another measure. Instead, Chi-square difference tests indicate that the data fit the hypothesized 5-factor model (with the items measuring the propensity to

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2  
3 morally disengage restricted to their own latent factor) ( $\chi^2 = 1551$ ,  $df =$   
4 979) significantly better ( $p < .001$  for all comparisons) than any of the  
5 models where the moral disengagement items were forced to load with  
6 Machiavellianism ( $\chi^2 = 1800$ ,  $df = 983$ ), moral identity ( $\chi^2 = 1723$ ,  
7  $df = 983$ ), perspective taking ( $\chi^2 = 1699$ ,  $df = 983$ ), or EC ( $\chi^2 = 1650$ ,  
8  $df = 983$ ). These results provide further evidence that the new propensity  
9 to morally disengage scale is empirically distinct from related constructs.

10 We therefore concluded from Study 1 that the propensity to morally  
11 disengage can be appropriately assessed with the new 8-item unidimen-  
12 sional scale. Appendix A presents these eight items (in bold), as well as  
13 the other sixteen items used to form the 16- and 24-item versions also  
14 considered in Study 1. In Studies 2–5, we continued the assessment of  
15 the 8-item version’s psychometric properties, with particular focus on its  
16 distinctness from and incremental validity beyond additional constructs  
17 in its nomological network. Specifically, we used four separate samples to  
18 test whether the new 8-item version of the propensity to morally disengage  
19 scale predicts multiple types of unethical decisions after controlling for  
20 sets of measures of constructs from each of the three nomological network  
21 categories previously discussed, as well as a longer extant measure of the  
22 propensity to morally disengage.  
23

## 24 Incremental Predictive Validity Studies

### 25 *Study 2*

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27  
28 Study 2 tested Hypothesis 1—that the propensity to morally disen-  
29 geage will have incremental validity in the prediction of unethical decision  
30 making, after controlling for four morally relevant individual trait differ-  
31 ences (Machiavellianism, moral identity and two facets of empathy) and  
32 socially desirable response tendencies. The dependent variable was a self-  
33 report measure of several types of unethical behavior (cheating, lying, and  
34 stealing).  
35

### 36 *Methods and Measures*

37  
38  
39 *Participants and Procedure.* For extra credit, 272 students at a uni-  
40 versity in the Northeastern U.S. completed one survey (with the propen-  
41 sity to morally disengage and other independent variable measures) mid-  
42 way through an undergraduate business course (not an ethics course);  
43 242 of these students completed a second survey (including the unethi-  
44 cal behavior measure) 1-month later. Demographic data are presented in  
45 Table 1.

*The Propensity to Morally Disengage.* At Time 1, participants completed the 8-item propensity to morally disengage scale ( $\alpha = .76$  in this sample). Fit indices from a CFA indicate a good fit of the data to a 1-factor model (e.g., CFI = 1.0, NNFI = 1.0, RMSEA = 0, SRMR = .03).

*Control Variables.* As in Study 1, Machiavellianism was measured using the 20-item Mach IV (Christie, 1970,  $\alpha = .69$  in this sample), moral identity was measured using Aquino and Reed's (2002) 5-item internalization measure ( $\alpha = .84$  in this sample), and perspective taking and EC were measured using the two seven-item subscales from Davis' (1983) IRI ( $\alpha$ s = .75 and .78, respectively, in this sample). All participants also completed the same 10-item measure of social desirability as in Study 1 (Strahan & Gerbasi, 1972).

*Dependent Variable.* We sought to first test whether the new propensity to morally disengage measure would show incremental validity in the prediction of a broad range of unethical behavior, and thus used Detert and colleagues' (2008) measure of cheating, lying, and stealing. This measure, obtained 1-month after the initial survey, asks respondents to indicate how often they have engaged in 13 unethical behaviors (e.g., "Taking low-cost items from a retail store"), using a 5-point scale ranging from "never" to "many times" ( $\alpha = .82$ ).

### Results

Means, standard deviations, alpha reliabilities (where appropriate), and bivariate correlations for the Study 2 variables are reported in Table 4. The correlations between the propensity to morally disengage scale and other measures are consistent with expectations, and with the results from Study 1. In this sample as well, the propensity to morally disengage correlates positively with Machiavellianism ( $r = .46, p < .01$ ), and negatively with moral identity ( $r = -.42, p < .01$ ) and two facets of empathy: perspective taking ( $r = -.33, p < .01$ ) and EC ( $r = -.48, p < .01$ ). Again, none of these relationships are so strong as to indicate that the propensity to morally disengage is redundant with any of the other constructs. We further confirmed this by estimating and comparing the results of five CFA models: one where the propensity to morally disengage items and those tapping the other four nomological network variables are all represented by their own latent factors ( $\chi^2 = 1950, df = 1024$ ), and four that force the propensity to morally disengage items to load with Machiavellianism ( $\chi^2 = 2010, df = 1029$ ), moral identity ( $\chi^2 = 2233, df = 1029$ ), perspective taking ( $\chi^2 = 2290, df = 1029$ ), or EC ( $\chi^2 = 2152, df = 1029$ ), respectively. Chi-square difference tests confirm that, in every case, the

TABLE 4  
Means, Standard Deviations, Correlations among Study 2 Variables

Variable	<i>M</i>	<i>SD</i>	1	2	3	4	5	6	7
1. Propensity to morally disengage <sup>a</sup>	2.79	.87	(.76)						
2. Machiavellianism <sup>a</sup>	2.77	.41	.46**	(.69)					
3. Moral identity <sup>a</sup>	6.29	.92	-.42**	-.44**	(.84)				
4. Perspective taking <sup>a</sup>	4.49	.98	-.33**	-.21**	.25**	(.81)			
5. EC <sup>a</sup>	5.02	.94	-.48**	-.43**	.55**	.35**	(.79)		
6. Social desirability <sup>a</sup>	3.61	1.33	-.16**	.07	-.04	-.07	-.07		
7. Self-reported cheating, lying and stealing <sup>b</sup>	1.83	.50	.31**	.20**	-.21**	-.18**	-.16*	-.08	(.82)

<sup>a</sup>*N* = 272 at Time 1, <sup>b</sup>*N* = 245 at Time 2. \*\**p* < .01. \**p* < .05. Where appropriate, alpha reliabilities appear along the diagonal.

data fit the model with a separate factor for the propensity to morally disengage significantly better ( $p < .001$  for all comparisons), providing further evidence that the propensity to morally disengage is empirically distinct from related constructs.

The positive and significant bivariate relationship ( $r = .31, p < .01$ ) with self-reported cheating, lying, and stealing behavior provides initial evidence of criterion validity for the new propensity to morally disengage scale. In fact, the propensity to morally disengage has the strongest bivariate relationship (among the Study 2 variables) with this measure of unethical behavior. To test Hypothesis 1, we examined the additional variance accounted for by the propensity to morally disengage in the prediction of self-reported cheating, lying and stealing, beyond that explained by four other morally relevant individual differences and social desirability. In the first step of a linear regression (see the first and second columns of Table 5), we entered the measures of Machiavellianism, moral identity, perspective taking, EC, and social desirability. In the second step, we added the propensity to morally disengage measure. In neither step are any of the control variables individually predictive of self-reported cheating, lying, and stealing. However, in support of Hypothesis 1, the propensity to morally disengage significantly predicts self-reported unethical behaviors ( $\beta = .22, p < .01$ ). Adding this variable to the equation increases  $R^2$  by 3% ( $\Delta F[1,236] = 7.96, p < .01$ ).

TABLE 5  
*Incremental Validity of the Propensity to Morally Disengage over Other Ethical Decision Making Predictors (Studies 2–4)*

	Study 2		Study 3		Study 4	
	Self-reported cheating, lying and stealing		Decision to intentionally misrepresent facts to client		Decision to protect one's self-interest over treating subordinates fairly at work	
	Model 1	Model 2	Model 1	Model 2	Model 1	Model 2
1. Machiavellianism	.11	.07				
Moral identity	-.14	-.10				
Perspective taking	-.13	-.09				
Empathetic concern	.00	.06				
Social desirability	-.08	-.04				
2. Propensity to morally disengage		.22**				
$R^2$	.08	.11				
$\Delta R^2$		.03**				
1. CMD			.99 (.97–1.01)	1.00 (.98–1.02)		
Idealism			.84 (.67–1.06)	.92 (.72–1.17)		
Relativism			1.00 (.78–1.28)	.86 (.68–1.14)		
Social desirability			.95 (.74–1.22)	.97 (.76–1.24)		
2. Propensity to morally disengage			2.15 (1.35–3.43)**			
<i>Model chi square (df)</i>			3.16 (4)	14.06 (5)		
$\Delta$ <i>Chi square (Δdf)</i>				10.90 (1)**		
1. Dispositional shame					1.75 (.68–4.53)	1.41 (.534–3.74)
Dispositional guilt					.17 (.06–.47)**	.29 (.09–.86)*
Social desirability					.73 (.53–1.00)	.71 (.52–.98)*
2. Propensity to morally disengage					14.21 (3)	1.93 (1.13–3.31)*
<i>Model chi square (df)</i>					20.14 (4)	20.14 (4)
$\Delta$ <i>Chi square</i>						5.84 (1)*

Note. Odds ratios (Exp[B]) and 95% confidence intervals are presented for logistic regressions. Standardized coefficients are presented for the linear regression.  $N = 245$  for Study 2,  $N = 248$  for Study 3, and  $N = 225$  for Study 4. \*\*  $p < .01$ . \*  $p < .05$ .



*Study 3*

Study 3 tested Hypothesis 2—that the propensity to morally disengage will show incremental validity in the prediction of unethical behavior after controlling for three deliberative moral reasoning or ethical philosophy factors (cognitive moral development, idealism and relativism) and socially desirable response tendencies.

*Methods and Measures*

*Participants and Procedure.* Three hundred and four students (74% male,  $M_{\text{age}} = 28.5$  years,  $SD = 2.2$  years) in an international MBA program participated in the study. Participants had an average of 5.6 years of work experience prior to entering the program ( $SD = 1.9$ ) (see Table 1 for additional demographics). Respondents completed a survey with the 8-item propensity to morally disengage scale and other measures as part of their course participation in a class on ethics and corporate social responsibility. To minimize contamination of results due to the class' subject matter, they completed the survey before the class began. The measure of unethical behavior (the dependent variable) was collected as the class began (Time 2), as part of the first assignment.

*The Propensity to Morally Disengage.* We used the same 8-item propensity to morally disengage scale developed in Study 1 ( $\alpha = .70$  in this sample). The fit indices from a confirmatory factor analysis again revealed a good fit of the data to a 1-factor model (CFI = .99, NNFI = .99, RMSEA = .02, SRMR = .04).

*Control Variables.* Cognitive moral development was measured using the Defining Issues Test (DIT) (Rest, 1986, 1990). The original test presents respondents with six moral dilemmas, asking them to rank the four most important considerations they would use in making a decision about what to do in each dilemma. A score, which is computed from those considerations ranked highest, is considered a measure of the proportion of one's reasoning that is at the highest (principled) level of moral reasoning. We used a short form version of the DIT that uses three of the six scenarios; this version has been shown to correlate above .90 with scores on the longer version (Rest, 1990).

Idealism and relativism were measured using Forsyth's Ethics Position Questionnaire (1980), which measures an individual's level of agreement with general statements on a 9-point continuum, ranging from "completely disagree" to "completely agree." Ten items measured idealism ( $\alpha = .85$ , sample item: "It is never necessary to sacrifice the welfare of others") and ten items measured relativism ( $\alpha = .80$ , sample item: "What is ethical

TABLE 6  
*Means, Standard Deviations, Correlations Among Study 3 Variables*

Variable	<i>M</i>	<i>SD</i>	1	2	3	4	5
1. Propensity to morally disengage <sup>a</sup>	2.12	.77	(.70)				
2. Cognitive moral development <sup>a</sup>	40.44	16.88	-.23**				
3. Idealism <sup>a</sup>	5.84	1.40	-.27**	.01	(.85)		
4. Relativism <sup>a</sup>	4.82	1.31	.29**	-.10	-.18**	(.80)	
5. Social desirability <sup>a</sup>	3.49	1.34	-.05	-.03	.03	-.03	
6. Unethical decision (dummy variable coded 1 if decided to send fraudulent fax) <sup>b</sup>	.18		.23**	-.07	-.10	.01	.05

<sup>a</sup>*N* = 304 at Time 1, <sup>b</sup>*N* = 258 at Time 2. \*\**p* < .01. \**p* < .05. Where appropriate, alpha reliabilities appear along the diagonal.

varies from one situation and society to another”). All participants also completed the same social desirability measure used in Studies 1 and 2.

*Dependent Variable.* The unethical outcome used in this study was developed by coding MBA students’ responses to an ethically charged case assignment (“Conflict on a Trading Floor,” Badaracco & Useem, 2006) that asks participants to put themselves in the position of a junior trader at a large bank whose supervisor is requesting them to send a fax containing misleading and, in fact, fraudulent information to a client. Students were asked (prior to the beginning of the class) to write one page explaining what they would do if they were in the same situation and why. Responses ranged from sending the fax, to voicing concerns to others, to refusing to send the fax. These decisions about the case were coded by a trained independent coder. An unethical decision was represented by a dummy variable coded “1” if an individual explicitly stated that s/he would send the fax (18% of the sample stated this).

### Results

Means, standard deviations, alpha reliabilities (where appropriate), and correlations among the Study 3 variables are reported in Table 6. As expected, the propensity to morally disengage is modestly negatively correlated with cognitive moral development ( $r = -.23$ ,  $p < .01$ ) and idealism ( $r = -.27$ ,  $p < .01$ ), and positively correlated with relativism ( $r = .29$ ,  $p < .01$ ). Again, a CFA model where the moral propensity to morally disengage items are forced onto their own latent factor ( $\chi^2 = 827$ ,  $df = 347$ ) fit significantly better (both chi-square difference tests significant at  $p < .001$ ) than models forcing the propensity to morally disengage items to load on the same factor with the idealism items

( $\chi^2 = 2010$ ,  $df = 1029$ ) or the relativism items ( $\chi^2 = 2233$ ,  $df = 1029$ ). The computation of cognitive moral development makes it impossible to include it in CFA models. But, as the least strongly correlated with the propensity to morally disengage of the three nomological network variables included in this study (see Table 6), it is reasonable to conclude from the other results that these two constructs are also distinct.

As predicted, the propensity to morally disengage is positively correlated with the subsequently measured unethical decision to send the fax ( $r = .23$ ,  $p < .01$ ). In fact, the propensity to morally disengage is the only variable in this study that is significantly related to the dependent variable. To test Hypothesis 2, we examined the variance accounted for, beyond that explained by three moral reasoning and ethical philosophy factors and social desirability, by the propensity to morally disengage in the prediction of a work-related unethical decision. We used logistic regression to examine the difference between the chi-square statistics for the more and less restricted models (i.e., models that do and do not include the propensity to morally disengage). In the first step of the regression (see column 3 of Table 5), the decision to send the fax was regressed onto the measures of cognitive moral development, idealism, relativism, and social desirability. In the second step (column 4, Table 5), the propensity to morally disengage measure was added to the equation. The difference in the chi-square statistic between the first and second models is significant ( $\Delta\chi^2 = 10.90$  [1 d.f.],  $p < .01$ ). Adding the variable for the propensity to morally disengage ( $\text{Exp}[B] = 2.15$ ,  $p < .01$ ) in the second step contributes significantly to the prediction of whether an individual is likely to make the unethical decision. An odds ratio of 2.15 for moral the propensity to morally disengage means that, holding cognitive moral development, idealism, relativism, and social desirability constant, for every one unit (on the 1–7 scale) increase in an individual’s measured level of the propensity to morally disengage, s/he is 115% more likely to send the fraudulent fax. This result provides support for Hypothesis 2. Specifically, these results provide evidence that the propensity to morally disengage predicts an unethical decision beyond three important independent variables that represent the traditional rational/deliberative approach to ethical decision making.

#### *Study 4*

Study 4 tested Hypothesis 3—that the propensity to morally disengage will have incremental validity in the prediction of unethical behavior, measured here as a self-serving decision in a work context, after controlling for two measures of moral affect (dispositional shame and guilt) and socially desirable response tendencies.

### *Methods and Measures*

*Participants and Procedure.* For extra credit, 250 undergraduate students at a university in the Northeastern U.S. completed one survey (with the propensity to morally disengage and social desirability measures) midway through a business course (Time 1), and 225 of these students completed a second survey (containing the other measures, including the dependent variable) 1-month later (Time 2). Demographic data are presented in Table 1.

*The Propensity to Morally Disengage.* All participants completed the 8-item propensity to morally disengage measure at Time 1 ( $\alpha = .77$  in this sample). Fit indices from a confirmatory factor analysis again revealed a good fit of the data to a 1-factor model (CFI = .99, NNFI = .98, RMSEA = .04, SRMR = .04).

*Control Variables.* Dispositional shame and guilt were measured using the Test of Self-Conscious Affect (TOSCA; Tangney, Wagner, & Gramzow, 1989). The scale is comprised of 16 brief scenarios (for example, “You are driving down the road, and you hit a small animal”); after each scenario there are two statements, one tapping guilt (in this example, “You would feel bad if you hadn’t been more alert driving down the road”) and the other tapping shame (in this example, “You would think: ‘I’m terrible.’”). The mean of each set of eight statements (assessed on a 5-point scale ranging from “not very likely” to “very likely”) is considered a measure of respondents’ dispositional guilt ( $\alpha = .77$ ), and shame ( $\alpha = .70$ ), respectively. To control for socially desirable responding, all participants also completed the same 10-item measure of social desirability used in Studies 1–3.

*Dependent Variable.* The unethical outcome in this study was operationalized using responses to an exercise that allowed respondents to make a work-related self-serving decision. Participants read a one-page description of a managerial dilemma in which they are asked to play the role of a manager who discovers that a subordinate—though they do not know which one—has made an error that will cost the company \$90,000. The manager’s own supervisor then demands that s/he either personally take the blame (resulting in a loss of 50% of his/her own bonus) or blame a member of his/her team (costing the team member 50% of his or her bonus). The choice was coded as a self-serving decision if the respondent chose the latter option (11% made this choice), because the manager in the exercise cannot fairly blame any specific team member; doing so saves the manager’s (i.e., the respondent’s) bonus at the expense of his/her subordinate.

TABLE 7  
Means, standard deviations, correlations among Study 4 variables

Variable	<i>M</i>	<i>SD</i>	1	2	3	4
1. Propensity to morally disengage <sup>a</sup>	2.74	.86	(.77)			
2. Dispositional shame <sup>a</sup>	3.01	.51	.01	(.70)		
3. Dispositional guilt <sup>a</sup>	3.98	.46	-.47**	.31**	(.77)	
4. Social desirability <sup>a</sup>	3.69	1.34	.08	-.02	-.14*	
5. Self-interested decision (dummy variable coded 1) <sup>b</sup>	.11		.23**	.00	-.20**	-.09

<sup>a</sup>*N* = 250 at Time 1, <sup>b</sup>*N* = 225 at Time 2. \*\**p* < .01. \**p* < .05. Where appropriate, alpha reliabilities appear along the diagonal.

### Results

Means, standard deviations, alpha reliabilities (where appropriate), and correlations among the Study 4 variables are reported in Table 7. Bivariate relationships are consistent with expectations. The propensity to morally disengage correlates negatively with dispositional guilt ( $r = -.47$ ,  $p < .01$ ) and is uncorrelated with dispositional shame ( $r = .01$ ,  $ns$ ). As in Studies 2 and 3, CFA models forcing the propensity to morally disengage items to load onto their own latent factor ( $\chi^2 = 1641$ ,  $df = 737$ ) result in significantly better fit indices (both chi-square difference tests are significant at  $p < .001$ ) than models forcing the propensity to morally disengage items to load onto factors measuring shame ( $\chi^2 = 1797$ ,  $df = 739$ ) or guilt ( $\chi^2 = 1761$ ,  $df = 739$ ).

The positive relationship between the propensity to morally disengage and the self-serving decision at work ( $r = .23$ ,  $p < .01$ ) provides further evidence of criterion validity for the new propensity to morally disengage scale. Again, the propensity to morally disengage has the strongest bivariate relationship (among the independent variables) with the outcome measure. To test Hypothesis 3, we examined the additional variance accounted for, beyond that explained by the measures of moral affect and social desirability, by the propensity to morally disengage in the prediction of the self-interested work decision. We used logistic regression to examine the difference between the chi-square statistics for a more and less restricted model (i.e., models that do and do not include the propensity to morally disengage). In the first step of the regression (see the column 5 of Table 5), the self-serving decision was regressed onto dispositional shame, guilt, and social desirability. In the second step (column 6 of

Table 5), the measure of the propensity to morally disengage was added to the equation. The difference in the chi-square statistic between the first and second models ( $\Delta\chi^2 = 5.84$  [1 d.f.],  $p < .05$ ) is significant. Adding the propensity to morally disengage measure ( $\text{Exp}[B] = 1.93$ ,  $p < .05$ ) in the second step adds significantly to the prediction of the self-serving decision. An odds ratio of 1.93 for the propensity to morally disengage means that, holding shame, guilt, and social desirability constant, for each one unit increase in an individual's measured level of the propensity to morally disengage, s/he is 93% more likely to make a self-interested decision at the expense of a subordinate. This result provides support for Hypothesis 3.

#### Study 5

Study 5 tested Hypothesis 4—that the propensity to morally disengage will have incremental validity in the prediction of unethical behavior, measured here as the unethical work behaviors of current employees as rated by both their peers and supervisors, after controlling for the most relevant extant measure of the propensity to morally disengage.

#### *Methods and Measures*

*Participants and Procedure.* A total of 399 individuals—141 employees matched with 129 supervisors and 129 co-workers—participated in the study. The employees worked at a variety of different organizations in the Southeast U.S., including technology, government, education, insurance, legal, financial, manufacturing, food service, and retail organizations. They averaged 25.7 years in age, 3.1 years of experience in the organization, and 48.9% worked full time (see Table 1 for additional demographics). The employees' supervisors were 63.5% male and 67.7% Caucasian (18.1% Hispanic or Latino/a, 6.3% African-American, and 2.4% Asian American), with an average age of 37.5 years and an average organizational tenure of 8.8 years. The employees' co-workers were 42.2% male and 59.1% Caucasian (21.3% Hispanic or Latino, 7.9% African-American, and 7.1% Asian American), with an average age of 27.8 years and an average organizational tenure of 5.1 years.

We recruited participants using a snowball sampling procedure (e.g., Grant & Mayer, 2009; Skarlicki & Folger, 1997): Researchers sent an electronic message to 277 (51% response rate) students in upper-level management courses at a university in the Southeastern U.S. and provided them with the opportunity to partake in a study for extra credit. Students who worked at least 20 hours per week in a job were allowed to participate; otherwise, they were asked to invite a friend or family member

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2  
3 to participate. The “focal employees” were instructed to visit a website to  
4 complete a survey, and to send an electronic survey link to their supervi-  
5 sor and a co-worker familiar with their work. Respondents were assured  
6 that their responses would remain confidential. Focal employees provided  
7 demographic and other information as well as responses to items from  
8 two scales measuring the propensity to morally disengage. Co-workers  
9 and supervisors provided ratings for the unethical work behaviors of the  
10 focal employee.

11 *The Propensity to Morally Disengage.* Employees completed the 8-  
12 item propensity to morally disengage scale described in Study 1 ( $\alpha = .90$   
13 in this sample). Fit indices from a confirmatory factor analysis revealed  
14 an acceptable fit of the Study 5 data to a 1-factor model (CFI = .96,  
15 NNFI = .94, RMSEA = .12, SRMR = .05).

16 *Alternative Measure of Moral Disengagement.* We used Duffy and  
17 colleagues’ measure (Duffy et al., 2005; Duffy et al., 2002; McFerran  
18 et al., 2010) as an alternative measure of the propensity to morally dis-  
19 engage. It is the only extant measure that was designed to measure the  
20 propensity to morally disengage in adults (rather than children, as in Ban-  
21 dura et al., 1996; or undergraduate students, as in Detert et al., 2008),  
22 without being focused on a specific type of audience, such as athletes  
23 (Boardley & Kavussanu, 2007; Corrion et al., 2010), or computer hackers  
24 (Rogers, 2001). It is also the only measure of the propensity to morally  
25 disengage that taps this disposition in a workplace context, thus making  
26 it the strongest alternative to our measure. The measure has 15 items  
27 ( $\alpha = .90$ ).

28 *Social Desirability.* Participants also completed the same 10-item  
29 measure of social desirability as in Studies 1–4 (Strahan & Gerbasi, 1972),  
30 but in this study responses were measured on a 7-point scale ranging from  
31 “strongly disagree” to “strongly agree” ( $\alpha = .81$ ).

32 *Dependent Variable.* We collected data on employees’ unethical work  
33 behavior from two sources, their immediate supervisor and a co-worker.  
34 Because there is no validated survey measure of employee unethical  
35 behavior, we selected items from two extant measures of undesirable  
36 work behavior—Robinson and O’Leary-Kelly’s antisocial work behav-  
37 ior scale (1998) and Bennett and Robinson’s organizational deviance  
38 scale (2000)—that most clearly represent *unethical* organizational be-  
39 havior. Rather than relying solely on our own judgments, we obtained  
40 ratings from a sample of 29 individuals with graduate-level management  
41 training—including 16 faculty who teach and research in the area of  
42 business ethics. We asked them to rate how *unethical* each of the 26  
43 work behaviors on these two extant scales (seven from the antisocial  
44 behavior scale and 19 from the organizational deviance scale) is. These  
45 raters were told that although all of the 26 behaviors are clearly

undesirable, not all are necessarily unethical according to the definition of unethical behavior as actions that cause direct harm to another individual or that violate widely accepted moral norms in society. Respondents then rated—using a 10-point scale ranging from (1) “not unethical in the least” to (10) “extremely unethical”—how unethical they considered each of the 26 antisocial and deviance items. We then created a 7-item measure of “unethical work behavior” using the items that received an average rating of 8.0 or higher.<sup>2</sup> The mean “unethicality” of these seven items was 8.37 vs., by contrast, 5.33 for the 19 items not used.

The seven items selected for this measure of unethical work behavior are: “Falsifying a receipt to get reimbursed for more money than you spent on business expenses” ( $M = 9.31$ ,  $SD = 1.00$ ), “Discussing confidential company information with an unauthorized person” ( $M = 8.59$ ,  $SD = 1.84$ ), “Damaging property belonging to my employer” ( $M = 8.25$ ,  $SD = 1.76$ ), “Taking property from work without permission” ( $M = 8.24$ ,  $SD = 1.17$ ), “Saying or doing something to purposely hurt someone at work” ( $M = 8.17$ ,  $SD = 1.81$ ), “Using an illegal drug or consuming alcohol on the job” ( $M = 8.03$ ,  $SD = 2.01$ ), and “Making ethnic, religious, or racial remarks at work” ( $M = 8.03$ ,  $SD = 1.80$ ). These items indeed seem different from an ethics perspective compared with excluded items such as, “Spending too much time fantasizing or daydreaming instead of working” ( $M = 4.04$ ,  $SD = 1.86$ ), and “Griping with coworkers” ( $M = 3.07$ ,  $SD = 1.79$ ).

We then calculated, using the main sample, separate supervisor- ( $\alpha = .93$ ) and co-worker- ( $\alpha = .93$ ) rated employee unethical behavior scales by taking an average of that rating source’s responses to the seven items about the focal employee’s unethical work behavior. Supervisors and co-workers reporting the frequency with which they had observed the focal employee engage in each of the unethical behaviors, ranging from (1) “never” to (7) “very often.”

### Results

As expected, the two measures of the propensity to morally disengage are highly correlated ( $r = .77$ ,  $p < .01$ ). However, despite this, the data fit a 2-factor CFA model that loads the two measures’ items onto separate latent factors ( $\chi^2 = 518$ ,  $df = 225$ ) better (chi-square difference test significant at  $p < .001$ ) than a 1-factor model that forces all (23 of) the propensity to morally disengage items to load together ( $\chi^2 = 747$ ,  $df = 230$ ). The two

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<sup>2</sup>As a robustness check, we ran all the same analyses reported in this section using a second 11-item measure of unethical work behavior that instead uses a cutoff of 7.0 or higher for item inclusion. The results are substantively unchanged.



TABLE 8  
*Incremental validity of the propensity to morally disengage over other predictors of unethical work behavior (Study 5)*

	DV = Supervisor-reported employee unethical behavior		DV = Coworker-reported employee unethical behavior	
	Model 1	Model 2	Model 3	Model 4
	Social desirability	-.08	-.08	.06
Alternative moral disengagement measure	.43**	.22	.44**	.07
Propensity to morally disengage		.27*		.49**
$R^2$	.22	.25	.16	.26
$\Delta R^2$		.03*		.10**

Standardized coefficients are presented.  $N = 128$  (supervisors),  $N = 129$  (co-workers)  
 \*\* $p < .01$ . \* $p < .05$ .

measures correlate similarly with supervisor-reported unethical behavior ( $r = .47, p < .01$  for the new, shorter measure, and  $r = .46, p < .01$  for the Duffy et al. measure). However, the new, shorter measure of the propensity to morally disengage is more strongly correlated with co-worker reported unethical behavior ( $r = .52, p < .01$ ) than the alternative (Duffy et al.) measure ( $r = .41, p < .01$ ), although the alternative measure is more strongly correlated with social desirability ( $r = -.46, p < .01$ ) than our new measure ( $r = -.33, p < .01$ ).

To test Hypothesis 4, we examined the variance in supervisor- and co-worker-reported employee unethical behavior accounted for by our new propensity to morally disengage measure beyond that explained by the alternative measure of moral disengagement and socially desirable response tendencies. As shown in Model 1 of Table 8, the alternative measure of the propensity to morally disengage explains a significant amount of variance in employee unethical behavior when our new measure is absent. However, the coefficient for the alternative measure becomes non-significant (see Model 2, Table 8) when the new measure of the propensity to morally disengage is added to the equation. In support of Hypothesis 4, the variable for the new propensity to morally disengage scale ( $\beta = .27, p < .05$ ) explains an additional three percent of the variance in supervisor-rated employee unethical work behavior even after accounting for an alternative measure of the propensity to morally disengage.

This same pattern of results emerges when using co-worker ratings of unethical work behavior as the dependent variable, though here the incremental value of the new measure is much greater. As shown in Table 8, the alternative measure of the propensity to morally disengage is initially a significant predictor of co-worker rated employee unethical behavior (Model 3), but ceases to be when our new measure of the propensity to morally disengage is added to the equation (see model 4). Again in support of Hypothesis 4, the variable for the new propensity to morally disengage scale ( $\beta = .49, p < .01$ ) explains an additional ten percent of the variance in co-worker-rated employee unethical work behavior even after accounting for an alternative measure of the propensity to morally disengage. This finding is particularly impressive because co-workers may be in a significantly better position than supervisors to know about the various types of employee unethical behavior assessed here, as presumably employees are more motivated to hide these behaviors from supervisors.

In both Models 2 and 4 (in Table 8), the standardized coefficients for the new measure are larger than the coefficients for the alternative measure, another means of comparing the relative importance of the two measures in the prediction of unethical workplace behavior (Combs, 2010). Finally, when models (not presented here) are run including our measure of the propensity to morally disengage and social desirability in the first step, and the alternative measure in the second step, in neither case does adding the alternative measure explain a significant additional proportion of the variance in unethical behavior. Collectively, these results suggest that the new measure of the propensity to morally disengage is superior to the closest current alternative.<sup>3</sup>

### Discussion

The set of studies presented here demonstrates the potency of individuals' propensity to morally disengage in the prediction of unethical organizational behavior and, in the process, makes two major contributions to the literature. First, our finding—across five different samples and using different dependent variables—that the propensity to morally disengage consistently emerges as a significant predictor of a wide range of organizationally relevant unethical behaviors, explaining additional variance above and beyond many of the major alternative individual difference antecedents, provides compelling evidence that morally

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<sup>3</sup>We also ran all models described in Study 5 using as dependent variables the full-length scales for antisocial workplace behavior (Robinson & O'Leary-Kelly, 1998) and organizational and interpersonal deviance (Bennett & Robinson, 2000). Results are similar results to those reported here (and are available from the first author).

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2  
3 disengaged reasoning represents a critical factor in a wide range of uneth-  
4 ical workplace behaviors.

5 Specifically, we demonstrated the predictive validity of the propensity  
6 to morally disengage beyond the explanatory power of (1) morally salient  
7 individual differences (Machiavellianism, moral identity, and empathy),  
8 (2) rational-deliberative constructs such as capacity for moral reasoning  
9 (cognitive moral development) and philosophical orientations (predis-  
10 positions towards idealism and relativism), and (3) dispositional moral  
11 emotions (shame and guilt) across an array of unethical behaviors, includ-  
12 ing self-reported lying, cheating, and stealing, business-related unethical  
13 and self-serving decision making, and supervisor- and co-worker-reported  
14 unethical employee behavior. In fact, compared to the numerous alterna-  
15 tive predictors we selected and controlled for based on theoretical and  
16 meta-analytic evidence, our measure of the propensity to morally disen-  
17 gage is the strongest and most consistent predictor of multiple unethical  
18 organizational outcomes in the studies reported here.<sup>4</sup>

19 Notably, most of the relationships between the propensity to morally  
20 disengage and other constructs in its nomological network were assessed  
21 for the first time, thus additionally contributing new knowledge about  
22 the morally relevant individual difference correlates of the propensity to  
23 morally disengage. As expected, the propensity to morally disengage cor-  
24 relates positively with Machiavellianism and relativism, negatively with  
25 moral identity, empathy, cognitive moral development, idealism and dispo-  
26 sitional guilt, and is not significantly correlated with dispositional shame.  
27 Given these findings, we suggest that future theoretical models, empirical  
28 studies, and intervention efforts seeking to understand and reduce uneth-  
29 ical outcomes include the propensity to morally disengage as a valuable  
30 explanatory construct.

31 Our second major contribution is the development and validation of a  
32 parsimonious, adult-oriented, easily administered measure of the propen-  
33 sity to morally disengage. Despite increasing interest in this construct in  
34 recent years (Aquino et al., 2007; Detert et al., 2008; Duffy et al., 2005;  
35 Moore, 2008b), research has been hampered by the lack of a consistently  
36 used, short, reliable and clearly valid measure of the propensity to morally  
37

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38  
39 <sup>4</sup>Note that in designing our studies to assess the incremental validity of our propensity  
40 to morally disengage measure we chose the baseline controls (which are entered as a block  
41 in the regressions) on the basis of theory and meta-analytic evidence. As shown in the  
42 correlation tables for studies 2, 3, and 4 (Tables 4, 6, and 7), the majority these variables are  
43 significantly correlated at the bivariate level with the dependent variable in these studies.

44 Because the bivariate relationships between cognitive moral development, idealism and  
45 relativism and the unethical organizational outcome in Study 3 are not significant, the  
incremental validity of the propensity to morally disengage demonstrated in this study is  
more limited.

disengage. Given consistent calls to improve measurement in the field of behavioral ethics (Mayer et al., 2009; Tenbrunsel & Smith-Crowe, 2008; Treviño et al., 2006), we believe this contribution warrants special note. We followed rigorous best practices for new measure development (Clark & Watson, 1995; Hinkin, 1995; John & Benet-Martinez, 2000) to construct and validate—using multiple samples diverse in respondent background and demographics—a parsimonious and psychometrically sound measure of the propensity to morally disengage in adults that should prove useful in a broad spectrum of future research with adults. The measure is consistent with Bandura’s theory of moral disengagement as a general process and taps, in a unifactorial scale, the eight theorized mechanisms that sever the link between internalized moral standards and unethical behavior.

In our analyses, we were also able to show that the new measure outperforms the best available alternative, and much longer, measure of the construct (McFerran et al., 2010). We also consistently controlled for respondents’ tendency to respond in socially desirable ways, thereby ruling out potential concerns that our findings might be contaminated by social desirability bias. Further, to reduce concerns about respondent fatigue, common method bias, and uncertain direction of causality as much as possible, we collected data for Studies 2–4 using two surveys each, with time separation of up to 1-month (Ostroff, Kinicki, & Clark, 2002). And, in Study 5, we obtained ratings of unethical employee behavior from both supervisors and co-workers.

#### *Implications for Theory*

In the series of studies presented here, the propensity to morally disengage is a consistently stronger correlate of unethical decisions and behavior than a wide array of other theoretically relevant predictors, and consistently explains variance in unethical decisions beyond these other predictors in multivariate analyses. Theoretically, this suggests that Bandura’s self-regulatory approach, and in particular, his articulation of the cognitive means by which people’s self-regulation process can be thwarted, provides a strong and useful foundation for understanding and predicting unethical decisions and behavior in organizations. Thus, behavioral ethics researchers should consider how other individual difference constructs might fit within a social cognitive self-regulation framework. For example, future research is needed to better understand how moral traits and emotions, such as empathy and guilt, relate to self-regulatory processes and morally disengaged reasoning. Such affective dispositions may be part-and-parcel of anticipatory self-regulation that prevents morally disengaged reasoning from occurring and/or may take effect and thus counter the personal acceptance of such reasoning as it reaches

consciousness. Similarly, deliberative ethical reasoning strategies (as driven by cognitive moral development or idealism/relativism) will likely only be used if moral disengagement has not occurred before such deliberation actually begins. Though the actual process of moral disengagement remains poorly understood, it may well occur prior to a number of these other processes, thus pre-empting them and helping to explain why the propensity to morally disengage is consistently a stronger predictor of unethical decisions and behavior in our studies.

Other constructs within the nomological network of the propensity to morally disengage might also be further considered from a self-regulation perspective. Machiavellianism, for example, may reflect either insufficient internalization of widely shared moral standards or the internalization of counter-normative standards that often dominate decision making when two or more standards conflict. In the first “Wall Street” movie, Gordon Gekko (with, for example, his “greed is good” speech) appears to be a strong Machiavellian who has internalized and is driven by self-interest rather than concern for others or fairness standards that are expected to have force in normally socialized individuals. In such a case, the propensity to morally disengage may be used to readily override more pro-social self-regulatory standards.

The social cognitive self-regulatory perspective may also provide theoretical insights if applied to some of the newer thinking and research in behavioral ethics. For example, much attention has been paid recently to less deliberative and more impulsive or intuitive models of ethical decision making (Haidt, 2001; Reynolds, Leavitt, & DeCelles, 2010). Our focal construct—the propensity to morally disengage—seems consistent with and potentially valuable to research on these models because this way of thinking likely happens automatically and, to some extent, at a subconscious level. Individuals with higher propensities to morally disengage may in fact be those with the most developed “self-justifying systems” for rationalizing intuitively made decisions. All of the ideas proposed here must be further theoretically developed and tested. Our point is to suggest potential ways that behavioral ethics scholars might benefit from adopting this overarching theoretical framework and, where appropriate, our new scale for empirical testing of emergent hypotheses.

#### *Implications for Research*

Following Bandura, we conceptualized the propensity to morally disengage as an individual difference that represents a generalized cognitive orientation to the world. However, the propensity to morally disengage is not necessarily invariant across contexts or time (Paciello, Fida, Tramontano, Lupinetti, & Caprara, 2008). Consistent with his broader

social cognitive theory, and with most work on moral development (Bandura, 1986; Kohlberg, 1984), Bandura described moral disengagement as explicitly interactive, the result of the continued reciprocal influences of the individual, behavior, and the environment (Bandura, 2002). Thus, acknowledging that disengaging self-sanctions through morally disengaged reasoning may also be triggered by specific contextual factors (suggesting moral disengagement may also have a state instantiation) is consistent with Bandura's theory (1999, 2000). For example, specific moral disengagement mechanisms are most likely triggered by particular circumstances (e.g., displacement of responsibility in response to dominant authority figures). Researchers are beginning to investigate situationally motivated moral disengagement as triggered by unethical actions (e.g., Shu, Gino, & Bazerman, 2011), or other aspects of a situation such as how much personal gain or harm to others is involved (e.g., Treviño, Detert, Sweitzer, & Gephart, 2008). A next step will be to combine these approaches by simultaneously studying both dispositional and situational influences, and their possible interaction, on morally disengaged reasoning (see Cronbach & Snow, 1977 for ideas on constructing research designs that capture both individual and situational variation). Research on moral disengagement that examines both the person and the situation simultaneously would capture more directly the interactionist nature that dominant theories agree characterize our moral selves (Bandura, 1990b; Treviño, 1986). Such approaches would also be consistent with thinking about many other individual differences (Hoyle & Leary, 2009) and with theories describing how traits can be activated in particular contexts (Tett & Burnett, 2003; Tett & Guterman, 2000).

Overall, our theoretical approach and statistical results are consistent with Bandura's theorizing (Bandura, 1986, 1990a, 1990b, 1999, 2002) that the specific mechanisms that underlie the propensity to morally disengage represent a single higher-order construct. Nevertheless, future research may reveal that valid and reliable first-order measures can also be developed to tap the propensity to use specific moral disengagement mechanisms (e.g., attribution of blame) that would be useful for particular research contexts or questions. For example, it may be valuable to assess only diffusion of responsibility when studying teamwork or collective action, or just displacement of responsibility when studying unethical behavior stemming from the nature of authority relationships in work hierarchies. Or, researchers may wish to compare various specific moral disengagement mechanisms to see which predict the most unwanted behavior in particular contexts. Because we focused on moral disengagement as an overarching propensity, we did not set out to develop first-order mechanism scales, and we did not produce subscales that presently meet acceptable standards (e.g., reliabilities of .70 or higher). Nevertheless, we

list all 24 items tested in Study 1 in Appendix A as a potential starting point for others who wish to pursue this research avenue.

Finally, this study and others have provided a great deal of information about the relationships between the propensity to morally disengage and many types of unethical behavior relevant to organizations as reported by the self or by employees' supervisors and co-workers. The linking here, in Study 5, of the propensity to morally disengage with supervisor and co-worker reports of unethical work behavior is particularly impressive because it is likely a conservative test of what employees actually do (because many unethical behaviors are intentionally hidden, especially from authority figures). Nonetheless, future studies that link the propensity to morally disengage with objectively measured subsequent unethical behavior in actual work contexts would be particularly valuable.

#### *Implications for Practice*

As Combs notes, the relevance of an empirical effect found in research is often best understood in economic terms (2010). In our studies, the additional variance in an outcome measure accounted for by our propensity to morally disengage scale ranges from 3% to 10%. At first blush, these effect sizes may seem to be of questionable practical value. Yet, given the enormous costs of unethical behavior to organizations and societies, we suggest that these results are actually very important. For example, three percent of the estimated \$350 billion that the U.S. government loses every year to tax evasion (U.S. Internal Revenue Service, 2010) is \$10.5 billion dollars, and three percent of the \$2.9 trillion estimated to be lost annually to global fraud (Association of Certified Fraud Examiners, 2010) is \$87 billion dollars. Clearly these losses, and many others of similar size (e.g., those occurring as a result of employee or consumer theft), are anything but trivial.

The magnitude of the effect sizes found here for the propensity to morally disengage on unethical decision making relative to other dominant predictors in the behavioral ethics literature (noting that the effects for the propensity to morally disengage found here are *beyond* many other theoretically important influences) is also worth considering (Aguinis et al., 2010). In the Kish-Gephart et al. (2010) meta-analysis, the largest average effect size for any of the traditionally studied individual difference predictors of unethical behavior was for Machiavellianism at  $\rho = .27$  (across 11 studies and 2,290 individuals). If the effects from the four predictive validity studies in this paper are combined, the average corrected correlation between our new propensity to morally disengage and unethical outcomes (using the same calculation methods as used in the

Kish-Gephart analyses, cf. Hunter & Schmidt, 2004) is  $\rho = .36$ .<sup>5</sup> Though this effect size only aggregates the effects of four studies and 857 individuals, it may be cautiously interpreted as suggesting that the propensity to morally disengage is the strongest individual difference predictor of unethical behavior identified to date.

The measure we developed in this study is also practically useful. It is short, easy to complete, and appropriate for general adult samples. The measure's consistently low correlation with the most commonly used social desirability measure allays concerns that individuals administered this measure for job-related purposes will simply provide "expected" responses. Thus, this measure may be useful in lieu of the much longer (100-plus items) and laborious administrations of more expensive commercial integrity tests (cf., Berry, Sackett, & Wiemann, 2007). Administering the scale to all applicants but having those making hiring decisions blind to the results may allow organizations to determine over time how valuable this instrument is in predicting subsequent misbehavior and whether its use as a selection tool is warranted.

Organizations should also investigate whether morally disengaged thinking can be attenuated through intervention or training. The one study that examined moral disengagement longitudinally found evidence that individual levels of moral disengagement are malleable to external influences over time (Paciello et al., 2008), which suggests that moral disengagement—even among those more predisposed toward such reasoning—may well be receptive to training interventions. Such interventions could have practical implications for organizations that are interested in reducing the harm caused by morally disengaged thinking. For example, if moral disengagement operates in ways similar to other cognitive biases (an empirical question), it could be quite beneficial to train employees to be on the lookout for certain modes of thinking (e.g., "I have to do it because my boss said so") so that they can catch themselves or others before unethical behavior occurs (Bazerman & Tenbrunsel, 2011). One interesting question in this regard would be whether training focused on those with initially high propensities to morally disengage (who might learn how to lessen their own distorted thinking) or those with lower initial propensities (who might learn more easily to recognize it and intervene with others) is more effective in reducing the measurable harm done from moral disengagement.

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<sup>5</sup>We thank Jennifer Kish-Gephart for her assistance with this analysis. Note that the 95% confidence intervals around the uncorrected meta-analytic correlations between unethical outcomes and (a) the propensity to morally disengage (uncorrected mean  $r = .30$ , CI .19 to .41) and (b) Machiavellianism (uncorrected mean  $r = .22$ , CI .15 to .28) overlap, as would be expected given the theoretical and empirical relationship between the propensity to morally disengage and Machiavellianism.



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2  
3 Organizations might also consider instituting a variety of measures  
4 designed to reduce the likelihood that morally disengaged reasoning goes  
5 unchecked. For example, leaders can use formal and informal means to  
6 increase individual accountability, making the displacement or diffusion  
7 of responsibility and the attribution of blame onto others feel less valid  
8 as justifications for behavior even among those with higher propensities  
9 to morally disengage. Leaders can likewise clearly encourage the use  
10 of ethical language and discourage the acceptance of euphemisms that  
11 cloud judgments. And, leaders can make harm to other stakeholders more  
12 real to employees, so that dehumanization or blame for “bringing harm  
13 onto themselves” is less likely. Of course, only with careful research will  
14 we know whether these interventions are effective in minimizing moral  
15 disengagement in those most predisposed to do so, or whether ultimately  
16 organizations are better off relying on careful selection and early screening  
17 mechanisms.

### 18 19 20 *Conclusion*

21  
22 This paper’s five complementary studies robustly demonstrate how  
23 individuals’ propensity to morally disengage predicts multiple types of  
24 work-relevant unethical behavior. We offer a systematically developed,  
25 parsimonious, reliable and valid new measure for assessing the propen-  
26 sity to morally disengage so that scholars and practitioners can use it to  
27 better understand and reduce unethical behavior. Collectively, our find-  
28 ings establish the pervasive predictive power of the propensity to morally  
29 disengage, illuminating the importance of this construct.

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1 CELIA MOORE ET AL.

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2  
3 APPENDIX A

## 4 Propensity to Morally Disengage Scale

5  
6 Moral Justification7  
8 **It is okay to spread rumors to defend those you care about.\***

9 It is alright to lie to keep your friends out of trouble.\*

10 Playing dirty is sometimes necessary in order to achieve noble ends.

11  
12  
13 Euphemistic Labelling14  
15 **Taking something without the owner's permission is okay as long as**  
16 **you're just borrowing it.\***

17 It's okay to gloss over certain facts to make your point.\*

18 When you're negotiating for something you want, not telling the whole  
19 story is just part of the game.20  
21 Advantageous Comparison22  
23 **Considering the ways people grossly misrepresent themselves, it's**  
24 **hardly a sin to inflate your own credentials a bit.\***25 Compared to other illegal things people do, taking something small from  
26 a store without paying for it isn't worth worrying about.\*27 Damaging property is no big deal when you consider that others are  
28 assaulting people.29  
30 Displacement of Responsibility31  
32 **People shouldn't be held accountable for doing questionable things**  
33 **when they were just doing what an authority figure told them to do.\***34 People cannot be blamed for misbehaving if their friends pressured them  
35 to do it.\*36 You can't blame people for breaking the rules if that's what they were  
37 taught to do by their leaders.38  
39 Diffusion of Responsibility40  
41 **People can't be blamed for doing things that are technically wrong**  
42 **when all their friends are doing it too.\***43 It's okay to tell a lie if the group agrees that it's the best way to handle the  
44 situation.\*

45 In contexts where everyone cheats, there's no reason not to.

1 48

## PERSONNEL PSYCHOLOGY

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3 Distortion of Consequences

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5 **Taking personal credit for ideas that were not your own is no big**  
6 **deal.\***

7 Walking away from a store with some extra change doesn't cause any  
8 harm.\*

9 It is OK to tell small lies when negotiating because no one gets hurt.

10

11 Dehumanization

12

13 **Some people have to be treated roughly because they lack feelings**  
14 **that can be hurt.\***

15 It's okay to treat badly somebody who behaves like scum.\*

16 Violent criminals don't deserve to be treated like normal human  
17 beings.

18

19 Attribution of Blame

20

21 **People who get mistreated have usually done something to bring it on**  
22 **themselves.\***

23 If a business makes a billing mistake in your favor, it's okay not to tell  
24 them about it because it was their fault.\*

25 If people have their privacy violated, it's probably because they have not  
26 taken adequate precautions to protect it.

27 Items measured on a 7-point Likert scale ranging from "strongly disagree"  
28 to "strongly agree".

29 Items in bold comprise the final 8-item measure. Items marked with \*  
30 comprise the 16-item measure.

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### Queries

- Q1** Author: Tenbrunsel & Messick (2004) is not cited in reference list. Please check.
- Q2** Author: Please provide complete bibliographic details in Association of Certified Fraud Examiners (2010).